

# Ma140a Probability

## Lecture: Moments and Law of Large Numbers

Juan Pablo Vigneaux

**Caltech**

Department of Mathematics  
California Institute of Technology  
Pasadena, CA, USA  
vigneaux@caltech.edu

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## Moments

Let  $(\Omega, \mathcal{F}, P)$  be a probability triple and  $X : \Omega \rightarrow \mathbb{R}$  a random variable. We denote by  $\Lambda_X$  the law of  $X$ ; it is a measure on  $\mathbb{R}$ .

The  $p$ -th moment of  $X$  is

$$\mathbb{E}(X^p) := \int_{\Omega} X^p(\omega) dP(\omega) = \int_{\mathbb{R}} x^p d\Lambda_X(x). \quad (1)$$

When  $X^n$  is not integrable, one says that the moment does not exist.

Similarly, the *centered  $p$ -th moment* is

$$\mathbb{E}((X - \mathbb{E}X)^p) = \int_{\mathbb{R}} (x - \mathbb{E}X)^p d\Lambda_X(x). \quad (2)$$

$\mathbb{E}X$  is called the *mean* of  $X$  (or average).  $\mathbb{V}(X) := \mathbb{E}((X - \mathbb{E}X)^2)$  is the *variance* of  $X$ . Its square root  $\sigma = \sqrt{\mathbb{V}(X)}$  is called *standard deviation*: it's a measure of "spread" of the distribution of  $X$  (in the same units as the mean).

## $\mathcal{L}^p$ spaces

For  $1 \leq p < \infty$  we say that  $X \in \mathcal{L}^p \equiv \mathcal{L}^p(\Omega, \mathcal{F}, P)$  if  $X$  is a r.v. such that  $\mathbb{E}|X|^p < \infty$ . remark that  $\mathcal{L}^p$  is a vector space.

We define  $\|X\|_p = (\mathbb{E}|X|^p)^{1/p}$ . It is a seminorm on  $\mathcal{L}^p$ , which means that

- it is homogeneous: for all  $\lambda \in \mathbb{R}$  and  $X \in \mathcal{L}^p$ ,  $\|\lambda X\| = |\lambda| \|X\|$ .
- It satisfies the triangular inequality (Minkowski's inequality): for all  $X, Y \in \mathcal{L}^p$ ,

$$\|X + Y\|_p \leq \|X\|_p + \|Y\|_p.$$

(The case  $p = 2$  easily follows from the Cauchy-Schwarz inequality, see next slide.)

Remark that for any  $p \geq 1$ ,  $\|Y\|_p = 0$  iff  $Y = 0$  almost everywhere. Let  $\mathcal{N}_p$  denote the subset of  $\mathcal{L}^p$  of functions that vanish almost everywhere. It is a vector space and the quotient  $L^p = \mathcal{L}^p / \mathcal{N}_p$  is a normed vector space called *Lebesgue space*. One can prove that it's also complete (hence a Banach space).

# Cauchy-Schwarz inequality

## Theorem

If  $X, Y \in \mathcal{L}^2$ , then  $XY \in \mathcal{L}^1$  and

$$|\mathbb{E}(XY)| \leq \|X\|_2 \|Y\|_2.$$

## Proof.

Remark first  $|\mathbb{E}XY| \leq \mathbb{E}|XY|$ , hence we can restrict to  $X \geq 0$  and  $Y \geq 0$ . To ensure convergence of integrals, truncate first:  $X_n = X \wedge n = \min(X, n)$  and  $Y_n = Y \wedge n$ . For all  $a, b \in \mathbb{R}$ , one has

$$0 \leq \mathbb{E}(aX_n + bY_n)^2 = a^2\mathbb{E}X_n^2 + 2ab\mathbb{E}X_nY_n + b^2\mathbb{E}Y_n^2.$$

The resulting quadratic in  $a/b$  or  $b/a$  is nonnegative, and the desired inequality follows from considering the necessary nonpositivity of its discriminant:

$(2\mathbb{E}X_n\mathbb{E}Y_n)^2 \leq 4(\mathbb{E}X_n^2)(\mathbb{E}Y_n^2) \leq 4(\mathbb{E}X^2)(\mathbb{E}Y^2)$ . Use the monotone convergence theorem to conclude. □

## Monotonicity of $L^p$ -norms

### Theorem (Williams, 6.7)

Suppose  $1 \leq p \leq r < \infty$ . If  $Y \in \mathcal{L}^r$ , then  $Y \in \mathcal{L}^p$  and  $\|Y\|_p \leq \|Y\|_r$ .

### Proof.

For  $n \in \mathbb{N}$ , we truncate  $X$ : let  $X_n(\omega) = (|Y|(\omega) \wedge n)^p$ . Let  $c(x) = x^{r/p}$ . Because both  $X_n$  and  $c(X_n)$  are in  $\mathcal{L}^1$  (since  $X_n$  is bounded) and  $c$  is convex, Jensen's inequality implies that

$$(\mathbb{E}X_n)^{r/p} \leq \mathbb{E}X_n^{r/p} = \mathbb{E}((|Y| \wedge n)^r) \leq \|Y\|_r^r.$$

Then use the monotone convergence theorem to conclude. □

This truncation method is very important.

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## Independence means multiply

Theorem (cf. Williams 7.1; Tamuz 8.1)

*Let  $X, Y \in \mathcal{L}^1$  be independent. Then  $X \cdot Y \in \mathcal{L}^1$  and*

$$\mathbb{E}(X \cdot Y) = \mathbb{E}X \cdot \mathbb{E}Y.$$

*In particular, if  $X$  and  $Y$  are independent elements of  $\mathcal{L}^2$ , then*

$$\text{Cov}(X, Y) := \mathbb{E}((X - \mathbb{E}X)(Y - \mathbb{E}Y)) = 0 \quad \text{and} \quad \mathbb{V}(aX + bY) = a^2\mathbb{V}(X) + b^2\mathbb{V}(Y).$$

**Proof.**

Apply the standard machine: assuming  $X$  and  $Y$  positive, approximate them with an increasing sequence of simple functions. For indicators, equality follows from the definition of independence. □

Remark: for any measurable functions  $f, g : \mathbb{R} \rightarrow \mathbb{R}$ ,  $g(X)$  and  $g(Y)$  are independent and  $\mathbb{E}f(X)g(Y) = \mathbb{E}f(X)\mathbb{E}g(Y)$  will follow provided the integrability conditions hold.

## Strong law in $\mathcal{L}^4$

Theorem (cf. Williams 7.2, Tamuz 8.3)

*Suppose that  $X_1, X_2, \dots$  are independent random variables, all with mean  $\mu$ , and that there is a  $K \in [0, \infty)$  such that  $\mathbb{E}X_i^4 \leq K$  for all  $i \geq 1$ . Let  $S_n = X_1 + \dots + X_n$ . Then  $P(n^{-1}S_n \rightarrow \mu) = 1$ . In other words:  $S_n/n \rightarrow \mu$  almost surely.*

Proof.

Suppose first the variables are centered:  $\mathbb{E}X = 0$ . Expand  $\mathbb{E}S_n^4 = \mathbb{E}(X_1 + \dots + X_n)^4$  and use the “independence means multiply” theorem to justify that most terms vanish, except those of the form  $X_i^4$  and  $X_i^2 X_j^2$  (one also needs monotonicity of norms to justify that finite fourth moment implies finite third and second moment).

Conclude that  $\mathbb{E}S_n^4 \leq 3Kn^2$ , hence that  $\mathbb{E}(\sum_n (S_n/n)^4) < \infty$ . It follows that  $\sum_n (S_n/n)^4$  converges almost surely (a.s.) and therefore  $S_n/n \rightarrow 0$  a.s. □

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## More inequality

$(\Omega, \mathcal{F}, P)$  probability space,  $X$  r.v. defined on it.

### Markov's inequality

If  $X$  is positive, then  $P(X \geq c) \leq \frac{\mathbb{E}X}{c}$ .

Proof: lower bound  $X$  by the simple function  $cI_{\{X \geq c\}}$ , then take expectations.

### Chebyshev's inequality

$P(|X - \mathbb{E}X| \geq t) \leq \frac{\mathbb{V}(X)}{t^2}$ .

Proof: Apply Markov's inequality to  $|X - \mathbb{E}X|^2$ .

## Theorem (Weak law of large numbers, cf. Tamuz 9.1, also Williams 7.3)

Let  $X_1, X_2, \dots$  be a sequence of independent real random variables in  $\mathcal{L}^2$ , such that  $\mathbb{E}X_n = \mu$  and  $\mathbb{V}X_n \leq \sigma^2$  for all  $n$ . Set  $Y_n = \frac{1}{n} \sum_{i=1}^n X_i$ . Then for every  $\varepsilon > 0$  and  $n \in \mathbb{N}^*$ ,

$$P(|Y_n - \mu| > \varepsilon) \leq \frac{\sigma^2}{n\varepsilon}.$$

### Proof.

Follows from Chebyshev's inequality, since  $\mathbb{V}(Y_n) = \frac{1}{n^2} \sum_{i=1}^n \mathbb{V}(X_i) \leq \sigma^2/n$ . □

By a truncation argument, the theorem also holds without supposing finiteness of the second moment, cf. Tamuz's Thm. 9.4.