Workshop 4

August 3, 2024

Unless said otherwise, the random variables involved in each problem are supposed to be defined on a common sample space $(\Omega, \mathcal{F}, \mathbb{P})$, and $\mathbb{E}X$ means $\int_{\Omega} X d\mathbb{P}$.

1 Laws of large numbers

Exercise 1. 1. Let X and Y be real valued random variables with finite mean and variance. Show that if X and Y are uncorrelated,¹ then for all $a, b \in \mathbb{R}$,

$$
Var(aX + bY) = a^2 Var X + b^2 Var Y.
$$
 (1)

2. Let $(X_n)_{n\in\mathbb{N}}$ be uncorrelated random variables, not necessarily identically distributed, with finite mean and variance. Set $\mu_i = \mathbb{E}(X_i)$ and $\sigma_i^2 = \text{Var}(X_i)$ for each $i \in \mathbb{N}$. Prove that for any $a > 0$,

$$
\mathbb{P}\left(\left|\sum_{i=1}^{n} X_i - \sum_{i=1}^{n} \mu_i\right| \ge a\right) \le \frac{\sum_{i=1}^{n} \sigma_i^2}{a^2}.
$$
\n(2)

Deduce that if all X_n are identically distributed, then for all $\epsilon > 0$, $\mathbb{P}\left(\left|n^{-1}\sum_{i=1}^n X_i - \mu_1\right| \geq \epsilon\right) \to 0 \text{ as } n \to \infty.$

Exercise 2 (Weierstrass' approximation theorem, cf. Williams 7.4). Let f be a continuous function on [0, 1]. Prove that for any $\epsilon > 0$, there exists a polynomial B such that

$$
\sup_{x \in [0,1]} |B(x) - f(x)| < \epsilon. \tag{3}
$$

(Hint: Consider the polynomials $B_n(x) = \sum_{k=1}^n f(\frac{k}{n}) {n \choose k} x^k (1-x)^{n-k}$. Write them as an expectation over a random variable S, and then express \tilde{S} as a sum of random variables. Use the weak $law.$)

2 Cramér-Rao inequality

The purpose of this section is to prove an important result in statistics: the Cramér-Rao inequality $(4).$ $(4).$

Exercise 3. Let $(\rho_\theta)_{\theta \in \Theta}$ be collection of probability measures, indexed by an open set $\Theta \subset \mathbb{R}$. We suppose that each ρ_{θ} has a density $f(x; \theta)$ with respect to (w.r.t.) the Lebesgue measure, that f is differentiable w.r.t. θ and moreover that $A = \{x : f(x; \theta) = 0\}$ does not depend on θ . We denote by \mathbb{E}_{θ} the integration w.r.t. $\rho_{\theta}^{\otimes n}$ (for some *n* that depends on the context).

A collection $\mathbf{X} = (X_1, X_2, ..., X_n)$ of random variables is called a random sample of size n for the population ρ_{θ} if $X_1, ..., X_n$ are independent and each has law ρ_{θ} . We denote by $f_n(\mathbf{x}; \theta)$ their joint density w.r.t. the n-dimensional Lebesgue measure. An estimator (e.g. of the unknown parameter θ) is a function $W(\mathbf{X}) = W(X_1, ..., X_n)$.

Suppose that

¹This means that $\mathbb{E}(XY) = \mathbb{E}(X)\mathbb{E}(Y)$.

- 1. $\frac{d}{d\theta} \mathbb{E}_{\theta} W(\mathbf{X}) = \int_{\mathbb{R}^n} \frac{\partial}{\partial \theta} (W(\mathbf{x}) f_n(\mathbf{x}; \theta)) d\mathbf{x}$, and
- 2. $Var_{\theta} W(\mathbf{X}) < \infty$.

Prove that

- 1. $\mathbb{E}_{\theta}(\frac{\mathrm{d}}{\mathrm{d}\theta}\log f_n(\mathbf{X};\theta))=0,$
- 2. $\mathbb{E}_{\theta}\left((\frac{d}{d\theta}\log f_n(\mathbf{X};\theta))^2\right) = n\mathbb{E}_{\theta}\left((\frac{d}{d\theta}\log f(X;\theta))^2\right)$, and that

3.

$$
\operatorname{Var}_{\theta} W(\mathbf{X}) \ge \frac{\left(\frac{\mathrm{d}}{\mathrm{d}\theta} \mathbb{E}_{\theta} W(\mathbf{X})\right)^2}{n \mathbb{E}_{\theta} \left(\left(\frac{\mathrm{d}}{\mathrm{d}\theta} \log f(X; \theta)\right)^2\right)}\tag{4}
$$

In particular, if W is an unbiased estimator of θ i.e. $\mathbb{E}_{\theta}W(\mathbf{X}) = \theta$, then

$$
\text{Var}_{\theta} W(\mathbf{X}) \ge \frac{1}{n \mathbb{E}_{\theta} \left(\left(\frac{\mathrm{d}}{\mathrm{d}\theta} \log f(X; \theta) \right)^2 \right)}.
$$
\n(5)

3 Some additional problems

Exercise 4 (Differentiation under the integral sign, cf. Cohn Ex 2.4.6). Let (X, \mathcal{F}, μ) be a measure space, I an open subinterval of R, and $f: X \times I \to \mathbb{R}$ a function. Suppose that

- 1. for each $t \in I$, the function $x \mapsto f(x, t)$ is integrable,
- 2. for each $x \in X$, the function $t \mapsto f(x, t)$ is differentiable, and
- 3. there is an integrable function $g: X \to [0, \infty)$ such that for all $t, t_0 \in I$,

$$
\left|\frac{f(x,t)-f(x,t_0)}{t-t_0}\right| \le g(x).
$$

Prove that

$$
\frac{\mathrm{d}}{\mathrm{d}t} \int_{X} f(x, t) \, \mathrm{d}\mu = \int_{X} \frac{\partial}{\partial t} f(x, t) \, \mathrm{d}\mu. \tag{6}
$$

Exercise 5. 1. Let Z be a random variable. Choose X and Y appropriately in Hölder's inequality (see Section 6.13 in Williams) to show that

$$
g(t) := \log \mathbb{E}|Z^t| \tag{7}
$$

is a convex function on the interval^{[2](#page-0-0)} of values of $t \geq 1$ such that $\mathbb{E}|Z^t| < \infty$. Pay attention to the conditions under which Hölder's inequality is valid.

2. Deduce from the previous result that if $r_1, ..., r_n$ are real numbers such that $\mathbb{E}|X^{r_i}| < \infty$ for all $i \in \{1, ..., n\}$ and $s = n^{-1} \sum_{i=1}^{n} r_i$, then

$$
\left(\prod_{i=1}^{n} \mathbb{E}|X^{r_i}|\right)^{1/n} \ge \mathbb{E}|X^s|.\tag{8}
$$

(Hint: Use induction in n .)

²It is an interval because of the monotonicity of p-norms: if $\mathbb{E}|Z^t| < \infty$, then $\mathbb{E}|Z^r| < \infty$ for every r in [1, t].